1 Abstract

2	The purpose was to examine the relationship between the rate of neural excitation (rate of rise in the
3	electromyogram, EMG) and the rate of isometric force development (RFD) to determine whether
4	surface EMG measures can detect nonlinearity that is expected due to underlying motor unit discharge
5	behavior and the summation of progressively larger motor unit potentials throughout recruitment. Due
6	to interest in obtaining a change point, a bilinear model was hypothesized to provide the best fit of the
7	EMG-RFD relationship compared to a linear model, exponential model and bilinear fit to log-
8	transformed data. 21 young adult participants performed isometric dorsiflexion contractions to 40% of
9	their maximal voluntary contraction (MVC) force. Contractions were performed in RFD conditions
10	ranging from slow (20 %MVC/s) to fast (peak volitional rate). The Akaike Information Criterion
11	supported nonlinear models in 16 of the 21 participants with the greatest overall support for the bilinear
12	model (n=13). The bilinear models indicated a mean change point at 204 %MVC/s. The present data do
13	not identify the specific motor unit control mechanisms at play and the influence of amplitude
14	cancellation on the electromyogram must be carefully considered.
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25 **1. Introduction**

26 Examining the relationship between neuromuscular excitation (NE) and force production 27 provides a means to study topics such as impaired motor control (Chou et al., 2013; Jahanmiri-Nezhad et 28 al., 2014; Ng et al., 1997), the effects of exercise training (Van Cutsem et al., 1998) and neuromuscular 29 efficiency (Paquin and Power, 2018). In some instances, physical function is predicted more strongly by 30 the rate of force development (RFD) than the peak force achieved (Bento et al., 2010; Hazell et al., 2007). NE primarily determines RFD (Maffiuletti et al., 2016), and is quantifiable using 31 32 electromyography (EMG). EMG represents the electrical sum of active motor units (Robertson et al., 33 2004) and is primarily determined by motor unit (MU) recruitment and rate coding mechanisms of force 34 control (Kamen and Gabriel, 2010). While one must not over-interpret measures from surface EMG with 35 respect to MU behavior, some recognize that nonlinearities in the EMG-force relationship may reflect 36 "different motor unit pool activation strategies" and have demonstrated that parameters from a bilinear 37 fit of the EMG-force relationship can be sensitive to experimental manipulations such as contraction 38 history (Paquin and Power, 2018). 39 At the MU level, the relationship between the rate of increase in current applied to the 40 motoneuron and RFD is linear (Baldissera and Campadelli, 1977). This linearity is due to bilinear firing 41 behavior of the alpha motor neuron offsetting the nonlinear input-output transform of muscle which 42 mimics a low-pass filter (Baldissera et al., 1998; Partridge, 1965). The bilinear relationship between 43 input to the motor neuron and its response (i.e. firing rate) includes a primary range of firing rates typically observed during slow contractions and a secondary range of firing rates observed during rapid 44 45 contractions or movements (e.g. Harwood, Davidson, & Rice, 2011; Kernell, 1965b). The two linear 46 ranges intersect at a change point and the secondary range has a greater slope. Feline studies have 47 demonstrated that both rapid muscle contractions from rest and higher frequency sinusoidal force

modulations depend on brief instances of secondary range MU discharge rates (Baldissera et al., 1998b)

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49 and the bilinear relationship between movement velocity and MU firing rates has been successfully 50 documented in humans (Harwood et al., 2011).

51 The dynamics of MU recruitment may also contribute to possible nonlinearity in the NE-EMG 52 relationship since higher threshold MUs have greater electrophysiological sizes (Masakado et al., 1994) 53 and are more likely to be recruited earlier in a contraction as RFD increases (J E Desmedt and Godaux, 54 1977; Yoneda et al., 1986). In slow muscle contractions, the greatest NE occurs close to peak force, 55 whereas during fast muscle contractions the greatest NE occurs closer to force onset (Ricard et al., 56 2005). Thus, bilinearity in rate coding and the nonlinear summation of progressively larger MU action 57 potentials are both considered as the basis of the present hypothesis that a bilinear relationship 58 between neural excitation and RFD can be observed with surface EMG measures. A more complete 59 understanding of this relationship will benefit applications of electromyography to the study of 60 neuromuscular function during rapid movements in health, pathology, and performance.

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2. Methods 62

63 2.1 Participants

Twenty-one healthy young adults, ten females and eleven males, (mean \pm SD: age=21.7 \pm 2.7 64 65 years, body mass = 73.6 ± 20.2 kg, height= 1.7 ± 0.1 m, maximal grip strength = 39.5 ± 10.2 kg) 66 participated in this study. Nine participants self-reported as consistently participating in high-intensity 67 physical activity for at least the previous six months. All participants were university students and free of neurological impairment, lower body dysfunction, and recent (<6 months) lower extremity injuries. 68 69 All participants signed a university approved informed consent before beginning the study.

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71 2.2 Procedures

72	EMG and isometric force recordings were obtained during a single testing session. Participants
73	were seated on a custom wooden bench with the left foot fastened with an inelastic strap to a plate
74	affixed to a strain gauge force transducer (Model SM-100, Interface Force Inc., Scottsdale, AZ). Force
75	was amplified and low-pass filtered at 50Hz at the time of recording (Model SGA, Interface force Inc.,
76	Scottsdale, AZ). The skin above the belly of the tibialis anterior muscle was shaved, abraded, and
77	cleansed with ethyl alcohol. A pre-amplified double differential surface electrode was secured to skin
78	above the mid-belly region of the tibialis anterior muscle (MA-300, Motion Lab Systems, Baton Rouge,
79	LA). The surface electrodes were 12mm diameter medical grade stainless steel disks with a 17mm inter-
80	electrode distance. A 13x3 mm reference bar separated the sensors and a ground electrode was placed
81	on the lateral malleolus. Amplification ranged from 2000 to 5700. Input impedance for this system is
82	>100 M Ω with a common mode rejection ratio >100 dB at 65Hz and noise <1.2uV RMS. Signals were
83	digitized at 2kHz with 24-bit resolution (cDAQ-9178, module NI9239, National Instruments, Austin TX).
84	DASYLab v.13 (National Instruments, Austin, TX) was used to control data acquisition and to provide
85	real-time force biofeedback.

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87 2.3 Experimental Conditions

88 Participants performed three maximum voluntary isometric contractions (MVCs) with the 89 maximum force achieved used to present relative force levels (%MVC) in visual feedback. Participants 90 were asked to produce force to match static plots of different linear ramp force-time curves (figure 1). 91 There were five different ramp force RFD conditions (20 %MVC/s, 40 %MVC/s, 80 %MVC/s, 160 92 %MVC/s, and 200 %MVC/s) and one condition of rapid force pulses. All conditions were performed to 93 40 %MVC. Each condition was practiced and performed for multiple trials. Each ramp force within a 94 trial was separated by 2 seconds and rapid pulses by 1 second. Each trial contained six ramps with one 95 minute of rest between recordings. With feedback based on visual inspection by the investigator,

participants practiced each RFD ramp condition until five ramps of adequate performance were
obtained. To reduce order effects, the ramp conditions were counterbalanced across participants
followed by two trials of rapid force pulses. After the conservative exclusion of contractions that
exhibited poor performance (typically excessive RFD at the onset of a ramp, large corrections during a
ramp, or poor amplitude control in pulses) an average of 57 contractions were analyzed in each
individual.

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Figure 1. A sample force-trace for the ramp force-matching condition (top) and graphs showing details
of data analysis (middle and bottom). The top graph contains a static plot of the 40 %MVC/s ramp
condition (black line) and the force produced for the entire trial by the participant (gray line). The
middle graph is isolates a single ramp between from the top graph with the addition of the dF/dt (RFD,
dotted line). The bottom graph is the rectified (gray line), smoothed (black line), and dEMG/dt (RER,
dotted line) EMG from the same ramp

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111 2.4 Signal processing

112 Force and EMG data were processed using LabVIEW v. 2014 (National Instruments, Austin, TX). 113 All values derived from the force-time curve were normalized to MVC force. An RFD time series was 114 calculated from the force-time curve as the slope from a linear fit line of all data points within a .1s 115 moving window (± .05s around each data point). After adjusting for gain, removing DC offset, and 116 bandpass filtering between 10-990Hz, the EMG was absolute value rectified. Based on recent work 117 involving EMG and rapid contractions, peak rate of EMG rise (RER) was selected to quantify NE because 118 it had the greatest correlation with RFD among measures that do not require the determination of EMG 119 onset, which would have been impractical and highly variable in the slowest RFD conditions (Josephson 120 and Knight, 2018). Using the same .1s window size as RFD computation, RER was calculated as the slope of the rectified, filtered (zero-lag 4th order low-pass Butterworth, 20Hz cutoff) electromyogram. The 121 122 EMG recordings were normalized to the RMS amplitude of EMG in the maximal MVC trial (± .250s 123 window surrounding MVC) which was filtered similarly.

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125 2.5 Model Selection

126 Based on the evidence of bilinearity in neuromuscular function cited above and with interest in 127 observing a potential change point, a strict bilinear model of the data was our primary model of interest. 128 Two other models were tested using guidance from research on blood lactate concentration curves. 129 Beaver, Wasserman, and Whipp (1985) determined that the best bilinear fit for this relationship is 130 achieved with a log-log transformation. Later researchers suggested that exponential model was most 131 representative of the underlying physiology (Hughson et al., 1987). A linear relationship between 132 surface EMG measures of NE and RFD, establishing our fourth model. Therefore, the models tested in 133 the present study were linear, bilinear, log-log transformation, and exponential.

135 The referent model (model 1) is a strict linear relationship, which is defined as: 136 y = ax + b137 138 where 'a' is the slope of the line, 'x' is the peak RFD, and 'b' is the y-intercept. 139 140 Model 2 is based on a strict bilinear relationship and is defined as: 141 142 $y = \begin{cases} a_0 + a_1 x & \text{if } x \le x_0 \\ b_0 + b_1 x & \text{if } x > x_0 \end{cases}$ 143 144 where $x_0 = \frac{a_0 - b_0}{a_1 - b_1}$ 145 146 147 where 'y' represents the estimated peak rate of NE, 'x' represents the peak rate of force development, 'a₀' represents a constant of the first linear relationship, 'a₁' represents the slope of the first linear 148 149 relationship, 'b₀' represents a constant of the second linear relationship, 'b₁' represents the slope of the 150 second linear relationship, and x₀ is the change point where the two relationships intersect. 151 Model 3 is a bilinear fit following a log-log transformation. For this model, the log values were found for 152 153 both peak RFD and peak RER prior to fitting it into the same bilinear relationship listed above. 154 155 Model 4 is based on an exponential relationship. This relationship is defined as: 156 $y = ae^{(bx)} + c$ 157

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where 'y' represents the estimated peak rate of neuromuscular activation, 'x' represents the peak rate
of force development, 'a' is the y-intercept, 'b' is the growth factor, and 'c' is a constant.

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162 2.6 Data analysis

163 The data from each participant was fitted with each model, using a custom LabVIEW program 164 (National Instruments, Austin, TX) to adjust model parameters until the mean squared error (MSE) was 165 minimized. The corrected Akaike Information Criterion (AICc, explained below) was computed for each 166 model. According to information theory, the model with the lowest AICc is most likely to be the best 167 model. The Akaike Information Criterion accounts for models with more adjustable parameters tending 168 to have lower mean squared error, even when not the best model (Akaike, 1973; Katsanevakis, 2006). 169 The formula for AIC is

169 The formula for AIC is

170 $AIC = nlog(MSE) + 2K + n(1 + log(2\pi))$

171 where n is the number of data points and K is the number of fitted parameters. Note that K should

include one extra parameter for the hidden estimate of residual variance (Burnham & Anderson, 2002),

and therefore K=3 for the linear model, K=5 for the bilinear and log-log models, and K=4 for the

exponential model. The formula for AICc (which is AIC corrected for small sample size (Akaike, 1973;

175 Shono, 2000)) is:

- 176 AICc = AIC + 2K(K + 1)/(n K 1)
- 177 When the sample size, n, is large, AICc approaches AIC.

The normalized model likelihood (Akaike weight, *w*_i) is the probability that model i is the best model, among the considered models (Burnham et al., 2002; Wagenmakers and Farrell, 2004). Akaike weight, is calculated as:

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$$w_i = \frac{\exp(-0.5\Delta_i)}{\sum_{k=1}^4 \exp(-0.5\Delta_k)}$$

182 where Δ_i is the difference between AICc for model i and AICc for the best model for that set of data:

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 $\Delta_i = AICc_i - AICc_{hest}$

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185 **3. Results**

The mean dorsiflexion strength was 34.04 <u>+</u> 7.30 N-m. During the rapid contractions, the peak RFD observed ranged from 287 to 623 %MVC/s with a mean peak RFD of 446 %MVC/s. The mean absolute peak RFD was 149 ±34.2 N-m/s.

189 For aggregate data, an exponential line of best fit had the lowest AICc (16015) and $w_{i=}91.2\%$.

190 Considering the potential for aggregate data to hide individual differences in best fit, model testing was

- 191 performed on an individual level, an approach consistent with the individual computation of blood
- 192 lactate curves (Hughson et al., 1987) and serves an interest in computing bilinear regression parameters
- such as the change point for individual research participants.
- 194 Table 1 shows mean squared error (MSE), corrected AIC (AICc), AICc difference (Δ), and relative
- 195 model likelihood (w, in percent) for the four models, for participant 1, to illustrate their computation
- and relationships. In this participant, the bilinear model has the lowest AICc, and therefore has the
- 197 greatest likelihood of being the best model. The AICc difference, Δ , is zero for the model with the
- 198 lowest/best AICc. The relative likelihoods of the four models add up to 100% (figure 2).
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- **Table 1**. Detailed model comparison in participant 1. Mean square error (MSE), corrected AIC (AICc), AICc difference (Δ), and normalized model likelihood (w) for each model. n=42 data points for this participant.

Quantity	Linear	Bilinear	Log-Log	Expon.
M.S.E.	21369	11011	199270	13272
AICc	544.5	521.7	643.4	527.0
Δ	22.8	0.0	121.6	5.3
w	0%	93%	0%	7%

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Figure 2. Linear (top), bilinear (middle), and exponential (bottom) fit for participant 1. AICc, AICc delta, and Akaike weight of each fit are listed on each figure. AICc delta is the difference between the AICc of that particular model and the lowest AICc observed among the three. Akaike weight is the likelihood

- 210 (percent) of a particular model being the best fit for that dataset.
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213	Table 2 shows the AICc differences (Δ) and the model likelihoods (w) for each participant. The
214	data in Table 2 indicate that a linear fit was best in five of the twenty-one participants while a non-linear
215	fit was best in the remaining sixteen. A chi square test indicated a significant (X^2 =5.76, p=0.01)
216	departure from an equal distribution across linear and nonlinear models. More specifically, linear model
217	was the strongest for five participants and had a better-than-5% chance of being the best model in two
218	other participants. The bilinear model was the strongest fit for thirteen participants and had a better-
219	than-5% chance of being the best model for the remaining eight participants. Log-log transformation
220	was the strongest fit for no participants and had a better-than-5% chance of being the best model in one
221	participant. Exponential was the strongest fit for three participants and had a better-than-5% chance of
222	being the best model in twelve additional participants.
223	Each bilinear fit has a change point: X-coordinate separating the primary range from the
224	secondary range. Table 3 provides the primary slope, change point, and secondary slope for each
225	participant in whom the bilinear model was most likely along with the coefficient of variation for each
226	parameter. The mean primary range slope was 0.51, the mean secondary range slope was 3.21, and the
227	mean peak RFD where NE changed from primary to secondary range was 204 %MVC/s. The change
228	point exhibited the least coefficient of variation.
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236 **Table 2**. Model comparison in all participants. Table shows AICc difference (Δ) and normalized model

237 likelihood (w, in percent) for each model in each participant. Bold indicates most likely model for each ls.

238	participant,	among the	tested	model

	Lin	ear	Bilinear		Log	Log-Log		Expon.	
Participant	Δ	w	Δ	w	Δ	w	Δ	w	
1	22.8	0%	0.0	93%	121.6	0%	5.3	7%	
2	6.4	2%	0.0	55%	1.2	30%	2.8	13%	
3	10.8	0%	0.0	78%	136.3	0%	2.5	22%	
4	10.7	0%	0.0	97%	93.7	0%	7.7	2%	
5	30.4	0%	2.6	22%	106.4	0%	0.0	78%	
6	0.0	78%	4.0	10%	110.8	0%	3.9	11%	
7	5.5	5%	0.0	80%	126.0	0%	3.3	15%	
8	6.9	2%	0.4	44%	113.8	0%	0.0	55%	
9	8.5	1%	0.0	98%	129.7	0%	10.7	0%	
10	4.3	8%	0.0	71%	120.2	0%	2.4	21%	
11	14.2	0%	0.0	96%	176.0	0%	6.4	4%	
12	0.0	61%	3.7	10%	69.5	0%	1.5	29%	
13	0.0	64%	1.1	36%	169.5	0%	17.3	0%	
14	5.7	3%	0.0	57%	116.3	0%	0.7	40%	
15	0.5	30%	0.3	33%	72.4	0%	0.0	38%	
16	0.0	49%	1.1	29%	131.6	0%	1.6	22%	
17	21.2	0%	0.0	86%	162.4	0%	3.7	14%	
18	0.0	71%	4.4	8%	67.9	0%	2.4	21%	
19	22.3	0%	0.0	98%	132.1	0%	7.4	2%	
20	5.6	5%	0.0	81%	132.8	0%	3.5	14%	
21	16.2	0%	0.0	100%	174.2	0%	10.9	0%	

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240 4. Discussion

241 This study sought to add to the current understanding of the rate of neural excitation (EMG rate 242 of rise, RER) throughout a wide range of isometric contraction rates. The aim was to determine whether there is support for a bilinear model of the NE-RFD relationship, considering the known bilinearity in MU 243 244 discharge behavior (Baldissera et al., 1998; Harwood et al., 2011; Kernell, 1965b) and the nonlinear 245 summation of MU potentials in the electromyogram as larger MUs are recruited (Masakado et al., 246 1994). Although we borrow the terms primary range and secondary range from studies that observed

- 247 bilinearity in MU firing rates, we do not suggest the observed bilinearity in the surface electromyogram
- is due specifically to this MU control mechanism.
- Table 3. Bilinear fit results. Primary range slope, change point, and secondary range slope, in each
- 250 participant for whom the bilinear model was most likely. A paired t-test revealed a significant difference
- 251 between the primary and secondary slopes (t=-6.67, p<0.001).

Subject ID Primary Range		Change Point	Secondary Range
	Slope	(RFD%MVC)	Slope
1	0.499	285.1	3.624
2	0.196	162.1	0.894
3	0.737	232.0	2.941
4	-0.174	154.0	1.954
7	-0.795	154.7	1.858
9	-0.186	120.5	1.793
10	0.102	126.3	1.796
11	1.334	256.7	6.024
14	1.249	222.0	3.986
17	0.675	181.8	2.063
19	1.611	367.3	7.994
20	1.692	256.8	4.386
21	-0.279	131.9	2.459
Mean	0.512	203.9	3.213
Standard Deviation	0.788	73.8	1.994
Coefficient of Variation (%)	154	36	62

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The AICc provided objective support for the bilinear model compared to linear, log-log, and exponential alternatives. While nonlinearity was not the best fit model in all participants, that 76% of the participants demonstrated a nonlinear best fit and all participants had a better-than-5% chance specifically for bilinearity supports the application of this model to the study of NE across increasing rates of force development. Bilinear model parameters from the 13 best-fit participants provided slopes of the primary and secondary excitation ranges and a change point. Among these three values the change point (203.9 ± 73.8 %MVC/s, coefficient of variation (CV)=36%) had the least variance across

261	participants followed by the secondary range slope (3.21 \pm 1.99, CV=62%). The slope of the relationship
262	between RER and RFD in the primary range was the most variable (.512 \pm .788, CV=154%).
263	While recognizing the limitations of surface electromyography to determine underlying MU
264	activity, one can still consider the possible contributions of rate coding and recruitment to the observed
265	bilinearity in EMG rate of rise. Specifically, one would expect a bilinear or exponential increase in EMG
266	as greater rates of descending excitation elicit secondary range firing rates (Baldissera et al., 1998;
267	Harwood et al., 2011) and/or recruit larger, high threshold motor units with larger electrophysiological
268	potentials (Masakado et al., 1994; Stalberg, 1980). The main challenge to this expectation is the
269	influence of amplitude cancellation (Keenan et al., 2005) in which the electromyogram is increasingly
270	attenuated at greater levels of excitation due to the summation of negative and positive phases of MU
271	action potentials. Since the effect of amplitude cancellation is more pronounced at greater levels of
272	excitation, the present findings of bilinearity in the relationship between EMG rate of rise and RFD might
273	be considered a possible underestimation of its true nature.
274	Determining why most, but not all, individuals had a nonlinear RFD-NE relationship requires
275	further consideration. Exploratory analysis comparing linear to non-linear subsets of participants was
276	performed for sex, grip-strength normalized-to-body mass, dorsiflexion MVC normalized-to-body mass,
277	BMI, body mass, and participation of high intensity activity in the previous year. Due to the small
278	sample size (N=21), Fisher's Exact Test was used for the influence of sex and activity. Independent t-
279	tests comparing linear and non-linear groups were used for BMI, body mass, normalized TA-MVC, and
280	normalized handgrip. The Fisher's Exact test revealed no differences in best fit by sex (p=0.635) or
281	regular participation in high intensity activity (p=0.611). No differences in BMI (t=0.853, p=0.440), body
282	mass (t=1.068, p=0.342), dorsiflexion MVC normalized to body mass (t=0.737, p=0.470), or normalized
283	handgrip strength (t=-0.424, p=0.677) existed.

284 As no significant differences of best-fit based on demographics or descriptive information arose, 285 other options should be considered. One must not only consider possible differences in MU 286 morphology, rate coding, and recruitment, but also the possibility of individual differences in factors 287 that contribute to amplitude cancellation (Keenan et al., 2005). Another possible explanation for the 288 mixed observations of nonlinearity across participants is heterogeneous compliance of the muscle 289 tendon (M-T) unit. As reviewed by Maffiuletti et al., the rate of force transmission through tissue is 290 partly determined by its stiffness and tendon stiffness in the lower extremity is known to be highly 291 variable across individuals (Maffiuletti et al., 2016). Some of the earliest published work on this topic 292 considered the manner in which pairs of electrical stimuli with brief intervals interact with tissue 293 compliance (e.g. Hill, 1949) and perhaps individuals with greater M-T unit stiffness might depend less on 294 secondary range MU firing rates during rapid contractions from rest, compared to individuals with less 295 M-T stiffness.

296 While extrapolation of specific MU control mechanisms from the surface electromyogram is not 297 recommended (Farina et al., 2014) this observation of nonlinearity in the EMG-RFD relationship is 298 consistent with expectations based on known nonlinearities in both MU rate coding and the summation 299 of progressively larger electrical potentials from higher threshold MUs. However, bilinearity was 300 observed in the MU firing rates of a study examining dynamic elbow extension across multiple angular 301 velocities, but not in surface EMG measures (Harwood et al., 2011). Although one could suggest that 302 differences in the EMG measures used might explain this discrepancy, the isometric equivalent of the 303 measure used by Harwood et al. (RMS amplitude from EMG onset to peak RFD), has a similarly strong 304 correlation with RFD as the rate of EMG rise measure used here (Josephson and Knight, 2018). One 305 could also suggest that surface EMG is more sensitive to recruitment than firing rate (Harwood et al., 306 2011) but such speculation seems to be based on publications that used slower 10 %MVC/s ramp 307 conditions which are less likely to elicit secondary range firing rates (Christie et al., 2009). It is possible

that the present study had greater sensitivity to detect bilinearity due to a greater number ofobservations used in model fitting at the level of the individual.

310 Different models have demonstrated the necessity of rapid initial MU firing rates to accomplish 311 rapid contractions (Baldissera et al., 1998; Del Vecchio et al., 2019; John E Desmedt and Godaux, 1977; 312 Heller, 2010) and found a lower RFD, decreased force, and a force lag when high initial MU firing rates 313 are removed. Considering the importance of RFD in mobility (Bento et al., 2010) and its responsiveness 314 to exercise training (Aagaard et al., 2002), knowledge of an EMG-RFD (or EMG-movement velocity) 315 change point may be informative in the practice of neuromuscular rehabilitation. Variance in the 316 location of the change point in our participants suggests that there may be an individual-specific 317 threshold above which the nonlinearities in recruitment or rate coding are expressed. In addition to 318 differences in M-T stiffness discussed above, it might be the case with humans in vivo that the change 319 point will also be influenced by muscle fiber length and contractile velocity. The observed variance in 320 the change point supports the value of examining bilinearity in individual participant data rather than in 321 group data.

322 As hypothesized, objective quantitative methods provided the greatest support for a bilinear 323 model of the EMG-RFD relationship, despite the known effects of amplitude cancellation which would 324 make such a finding less likely. We consider this finding to be preliminary and one that requires 325 replication as it has not been observed in other related experiments (Harwood et al., 2011) and the 326 results may be dependent on details of experimental design. Two known limitations should be 327 addressed in future studies. First, experimental conditions that produce more data points in the range 328 of the change point may enhance resolution. Second, extending the RFD conditions further into the 329 secondary range by performing force pulses to greater amplitudes would make quantification of the 330 EMG-RFD relationship more complete. Despite the limitations of surface electromyography, a more 331 complete understanding of the relationship between rates of neuromuscular activation and rate of force

332	development will improve our understanding of the neural control of rapid movement in health and

disease.

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